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# Fine Particulate Matter and Parkinson Disease Risk Among Medicare Beneficiaries

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## Abstract

### Background and Objectives

Numerous studies suggest that environmental exposures play a critical role in Parkinson disease (PD) pathogenesis, and large, population-based studies have the potential to advance substantially the identification of novel PD risk factors. We sought to study the nationwide geographic relationship between PD and air pollution, specifically $PM_{2.5}$ (particulate matter with a diameter <2.5 micrometers), using population-based US Medicare data.

### Methods

We conducted a population-based geographic study of Medicare beneficiaries aged 66–90 years geocoded to US counties and zip+4. We used integrated nested Laplace approximation to create age, sex, race, smoking, and health care utilization–adjusted relative risk (RR) at the county level for geographic analyses with $PM_{2.5}$ as the primary exposure of interest. We also performed an individual-level analysis using logistic regression with cases and controls with zip+4 centroid $PM_{2.5}$. We adjusted a priori for the same covariates and verified no confounding by indicators of socioeconomic status or neurologist density.

### Results

Among 21,639,190 Medicare beneficiaries, 89,390 had incident PD in 2009. There was a nationwide association between average annual $PM_{2.5}$ and PD risk whereby the RR of PD was 56% (95% CI 47%–66%) greater for those exposed to the median level of $PM_{2.5}$ compared with those with the lowest level of $PM_{2.5}$. This association was linear up to 13 $\mu g/m^3$ corresponding to a 4.2% (95% CI 3.7%–4.8%) greater risk of PD for each additional $\mu g/m^3$ of $PM_{2.5}$ ($p_{trend} < 0.0001$). We identified a region with high PD risk in the Mississippi-Ohio River Valley, where the risk of PD was 19% greater compared with the rest of the nation. The strongest association between $PM_{2.5}$ and PD was found in a region with low PD risk in the Rocky Mountains. $PM_{2.5}$ was also associated with PD in the Mississippi-Ohio River Valley where the association was relatively weaker, due to a possible ceiling effect at average annual $PM_{2.5}$ levels of ~13 $\mu g/m^3$.

### Discussion

State-of-the-art geographic analytic techniques revealed an association between $PM_{2.5}$ and PD that varied in strength by region. A deeper investigation into the specific subfractions of $PM_{2.5}$ may provide additional insight into regional variability in the $PM_{2.5}$-PD association.

## Introduction

Several studies have linked air pollution in the form of aerosolized particulate matter to various adverse health outcomes. Recent investigations identified associations between fine particulate matter, that is, particulate matter with diameter ≤2.5 micrometers ($PM_{2.5}$), and neurologic disease,

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## Glossary

**μg/m$^3$** = micrograms per cubic meter; **CI** = confidence interval; **GWR** = geographically weighted regression; **LISA** = local indicator of spatial association; **PD** = Parkinson disease; **$PM_{2.5}$** = particulate matter <2.5 micrometers in diameter; **R-INLA** = Laplace approximation method for Bayesian inference in R-project; **RR** = relative risk; **U.S.** = United States.

including dementia[1] and stroke.[2] The ultrafine particles (≤0.1 micrometers) in $PM_{2.5}$ cross the blood-brain barrier in humans.[3] In addition, some subcomponents of $PM_{2.5}$ are more neurotoxic than others. In particular, $PM_{2.5}$ can contain heavy metals, including arsenic and manganese, which have been implicated in the neuropathogenesis of basal ganglia degeneration.[4,5] Despite this, epidemiologic investigations of $PM_{2.5}$ and Parkinson disease (PD) have yielded mixed results,[6-20] with marked discrepancies in the magnitude, shape, and even direction of $PM_{2.5}$-PD associations. One possible contributor is the use of different PD-related outcomes and widely varying definitions of "incidence." In addition, the range of $PM_{2.5}$ levels across studies varies widely, and differences in the size and content of particulate matter in different regions might influence $PM_{2.5}$-PD associations. Accordingly, several prior studies found $PM_{2.5}$-PD associations to differ by geographic area[10-12] or urban/rural land use.[16,18] Several powerful spatial analytic methods offer to advance our understanding of the role of $PM_{2.5}$ in PD by enabling high-resolution, population-based investigations of the geographic distribution of PD, to characterize patterns of incidence and their relation to $PM_{2.5}$ exposure.

Similar to studies of $PM_{2.5}$ and PD, existing research into the national distribution of PD primarily consists of studies of mortality and prevalence.[21-23] One study found a North-South gradient in PD mortality and prevalence, which likely reflects the general burden of disease but does not capture the geographic patterns of *incidence* necessary for understanding the role of environmental risk factors. In addition, most nationwide studies of PD rely on state-level data.[22,24] To date, only one nationwide, county-level study of incident PD has been conducted in the United States.[21] This study found nonrandom clustering of PD in the Midwest and East South Central United States, which likely reflects, in part, the effect of environmental exposures on PD risk. Deeper exploration of these PD clusters, and the potential role of $PM_{2.5}$ in contributing to those clusters, requires further investigation with advanced geographic methods and geostatistical approaches not yet applied to neurodegenerative disease. We conducted a US population-based geographic study of PD risk to examine spatial patterns of newly diagnosed PD and relationships with $PM_{2.5}$, using a multimethod approach that included spatial analytic and statistical methods. We hypothesized that we would identify spatial clustering of PD and observe a positive association between $PM_{2.5}$ and PD, which would vary by region.

# Methods

This study was approved by the Human Research Protection Office at Washington University School of Medicine in St. Louis and by the Centers for Medicare & Medicaid Services. Participant consent in this records-based study was not required.

### Study Population and Case Ascertainment

Our eligibility criteria were designed to ensure a population-based sample with complete data: age-eligible for Medicare ≥2 years before diagnosis/selection (66 years and older), no Part C (Medicare Advantage/health maintenance organization) coverage, 90 years and younger, and US residence, all in 2009, without additional (e.g., medical) inclusion/exclusion criteria.[25] Incident PD cases included all study-eligible beneficiaries with at least one International Classification of Diseases, Ninth Revision, diagnosis code of 332 or 332.0 in 2009, but no prior year. This case definition, similar to a few prior studies of $PM_{2.5}$ and PD,[17,19,20] which maximizes sensitivity without materially affecting specificity,[26] aims to ensure representativeness of cases, including with regard to exposures. Beneficiaries with a diagnosis of atypical parkinsonism (333.0) or Lewy body dementia (331.82) were excluded if diagnosed in the year of PD diagnosis or earlier (4.6% potential cases).[25] We geocoded beneficiaries to their residential zip+4 after applying a two-year lag in their residential history, which we obtained from the 2007 Medicare beneficiary annual summary file. For maps and geographic analyses, we linked beneficiaries to county of residence to avoid zero inflation.

### $PM_{2.5}$ Exposure Estimation

Our exposure of a priori interest was average annual $PM_{2.5}$ from 1998 to 2000,[27] a period largely before PD onset. We used this period because it is likely etiologically relevant and 90% of beneficiaries maintained the same county code (84% had the same 5-digit zip code) for all years available (2004–2009). $PM_{2.5}$ data,[27] which were available in 1-kilometer grids, were based on several predictors, including satellite, meteorologic, land use, and elevation data with varying resolutions.[28] This $PM_{2.5}$ model achieved a cross-validated $R^2$ of 0.89. All 3 prior PD studies that used this or a similar $PM_{2.5}$ model[7,16,17] observed significant associations between $PM_{2.5}$ and PD, consistent with an acceptable degree of exposure measurement error in the context of air pollution health effects research, where associations are generally difficult to detect. We then used geographic information systems to estimate average annual $PM_{2.5}$ at 2 different geographic levels: (1) county for mapping and spatial analyses and (2) zip+4 (centroid) for individual-level regression analysis.

### Assessment of Covariates

For all essential covariate data, we started with data at the individual (beneficiary) level and only collapsed to the county level for mapping and spatial analyses. Covariates included beneficiary demographic information (age, sex, and race) and

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measures of health care utilization in the year before diagnosis/reference from the beneficiary annual summary file. We defined health care utilization as the number of physician visits (carrier) and outpatient visits either for the individual or for county-level analyses per county for all Medicare beneficiaries in our study. We obtained the county-level current prevalence of smoking cigarettes[29] for county-level/geographic analyses and estimated the probability of ever smoking at the individual level for individual-level analyses. We developed this smoking variable for use in geographic studies through a multivariable linear regression model with our validated claims-based probability of smoking[30] as the outcome. This gold-standard variable replicates the well-established relationship between smoking and PD.[25] Predictors of smoking probability in the new model were county-level prevalence of current smoking[29] and individual-level data from the beneficiary annual summary file. Similar to the original smoking model based on detailed claims, these individual-level predictors included sex, race, birth cohort, and selected medical conditions—here chronic obstructive pulmonary disease, lung cancer, stroke, acute myocardial infarction, other ischemic cardiovascular disease, stroke/transient ischemic attack, chronic kidney disease, osteoporosis, and depression. This smoking variable along with use of care, age, sex, and race represented our core set of covariates. As additional covariates for sensitivity analyses, we obtained the following ecologic data: census tract airborne trichloroethylene,[31] county-level area deprivation,[32] census block group median household income,[33] neurologists per county per 100,000 Medicare beneficiaries,[21] and county-level agricultural pesticide use.[34] In addition, from the beneficiary annual summary file, we obtained individual-level data on acute myocardial infarction, ischemic heart disease, stroke/transient ischemic attack, congestive heart failure, chronic obstructive pulmonary disease, and diabetes. These 6 conditions might be caused by $PM_{2.5}$ exposure and subsequently lead to care that facilitates diagnosis of PD, that is, act as nuisance mediators.

## Estimation of PD Relative Risk for Counties

We used integrated nested Laplace approximation for Bayesian inference in R-Project (R-INLA)[35] using R version 4.1.2 to estimate county-PD relative risks (RR) that account for known demographic risk factors[36-40] and spatial dependency. Specifically, these R-INLA–derived RRs were based on Gaussian distribution using indirect age-sex-race standardized incidence ratios, health care utilization, and smoking as input. We calculated the standardized incidence ratio by dividing the number of observed cases by expected counts and then multiplying by 100, based on 4 age (65–69, 70–74, 75–79, and 80+ years), sex, and 5 race (White, Black, Asian, Hispanic, and other) strata. To address spatial autocorrelation, we integrated spatial dependency into the R-INLA model using conditional autoregressive distribution to smooth risks according to the standardized incidence ratios of neighboring counties. The R-INLA–derived PD RR for each county was available for all county-level analyses.

## Assessment of Spatial Clustering of PD

To formally test for spatial clustering and identify high and low PD risk counties, we used univariate local indicators of spatial association (LISAs) to map PD hot and cold spots.[41] Hot-spot counties have above-average PD risk (RR) and share boundaries with counties that all have above-average PD risk. Cold-spot counties have below-average PD risk and share boundaries with counties that have below-average PD risk. LISAs identify where high and low-risk counties form contiguous clusters rather than only focusing on counties with significant RRs. LISAs also provide *Global Moran's I* value to describe the nationwide presence or absence of clustering.

## Examination of the $PM_{2.5}$-PD Association

We used 3 approaches to examine the relationship between $PM_{2.5}$ and PD—a regression model for assessing the *nationwide* relationship and 2 geographic approaches for assessing *regional* relationships. Specifically, for our nationwide assessment, we performed traditional multivariable regression at the individual level using Stata/MP version 17. For regional assessment, to explore whether the $PM_{.2.5}$-PD association differed by region, we used bivariate LISAs implemented in GeoDa version 1.20[41] and geographically weighted regression (GWR) implemented in ArcGIS. Bivariate LISAs were used to assess and visualize local spatial correlation between PD and $PM_{2.5}$ while GWR was used to determine and quantify the *direction and strength* of local associations.

### Logistic Regression Analysis

We performed logistic regression with PD as the outcome and zip+4 $PM_{2.5}$ as the independent variable, adjusted for the same a priori covariates as above but assessed at the individual level. To initially examine the association while allowing for nonlinear associations, we modeled $PM_{2.5}$ as deciles, with the lowest decile of $PM_{2.5}$ as the reference group. Based on these results, we then sought to develop a more parsimonious model using linear splines. We selected the final model using the Akaike information criterion while also examining the sensitivity of results to knot number and placement. To assess whether restriction to Medicare-aged beneficiaries could limit generalizability, we tested whether the $PM_{2.5}$-PD association differed by age. To assess whether associations might be due to occupational rather than environmental exposures, we tested whether the association was stronger in men than women. We tested for interaction on the multiplicative scale while including main-effects terms in the model. As a sensitivity analysis, we included Lewy body dementia cases. Because PD is relatively rare, the odds ratio provides an accurate estimate of the RR.

### Spatial Correlation

We used bivariate LISAs to overlay $PM_{2.5}$ and PD hot and cold spots and assess local spatial correlation.[41] A bivariate LISA map is the convergence of 2 univariate LISA maps into a single map (i.e., a map of PD hot and cold spots + a map of $PM_{2.5}$ hot and cold spots). A bivariate LISA map delineates 4

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**Table 1** Characteristics of Incident PD Cases and Controls With $PM_{2.5}$ Data Based on zip+4,[a] US Medicare 2009

| | PD cases (N = 65,180) | Controls (N = 15,561,435) |
|---|---|---|
| **Female, %[b]** | 51% | 58% |
| **Race/ethnicity, %** | | |
| **White** | 89% | 86% |
| **Black** | 6% | 8% |
| **Hispanic** | 2% | 2% |
| **Asian** | 2% | 2% |
| **Native American** | 0.1% | 0.2% |
| **Pacific Islander/other/unknown** | 1% | 2% |
| **Age, mean (standard deviation)** | 79 (6.1) | 76 (6.2) |

Abbreviations: PD = Parkinson disease; $PM_{2.5}$ = particulate matter <2.5 micrometers in diameter.
[a] All Medicare beneficiaries with incident PD or without PD in 2009 who met basic criteria; methods and sample are described in Silver et al. 2020.[43] Excludes 19,144 cases and 4,693,525 controls without geocodable zip+4 data for residence 2 years before PD diagnosis or the control reference date and 3,140 cases and 740,179 controls without $PM_{2.5}$ data at the geocoded location in the United States.
[b] Percentage excludes 47 controls with unknown sex.

cluster categories (high-high, low-low, low-high, and high-low). The high-high category represents counties where PD risk and $PM_{2.5}$ levels are both high relative to their means, and the low-low category represents counties where PD risk and $PM_{2.5}$ levels are both low; these categories are consistent with a positive correlation between $PM_{2.5}$ and PD. The 2 discordant (low-high and high-low) categories are inconsistent with a positive correlation.

### Spatial Regression

We used GWR to determine the *direction and strength* of the local associations between $PM_{2.5}$ and PD (ArcGIS). In GWR, local regression is performed for each county using surrounding counties. We used an adaptive spatial bandwidth to define the latter. GWR computed a regression coefficient for each county using county-PD RR as the dependent variable and $PM_{2.5}$ as the independent variable. We retained RR as a continuous outcome and modeled $PM_{2.5}$ in deciles in a single continuous variable. Thus, the GWR beta coefficients represent the absolute difference in the PD RR when going from any decile of $PM_{2.5}$ to the decile below. We then computed 95% confidence intervals (CIs) and mapped the GWR coefficients for counties where the CI excluded zero (significant at a two-sided alpha = 0.05). This allowed us to include both the positive and negative coefficients in a map, as well as to produce a map focused on regions where GWR coefficients were positive, providing a more conservative reflection of where the association between $PM_{2.5}$ and PD was positive. In addition, we used the Monte Carlo test of spatial variability to assess objectively whether the relationship between $PM_{2.5}$ and PD varied across the nation.

### Data Availability

The Centers for Medicare & Medicaid Services does not permit data sharing under the data use agreement.

## Results

After excluding prevalent PD, 21,639,190 US Medicare beneficiaries from Medicare research files[42] met initial study eligibility criteria,[43] including 89,790 PD incident cases and 21,549,400 non-cases. Of these, we had 65,180 cases (73%) and 15,561,435 non-cases (72%) with high-resolution residence information available (zip+4) 2 years before PD diagnosis or the control reference date. We were restricted to those with zip+4 information so that our ecologic exposure estimates would better represent individual-level exposures. We observed no demographic differences between the full data set[43] vs this sample (Table 1), overall or by case status. In both samples, 89% of cases and 86% of controls were non-Hispanic White and 6% of cases and 8% of controls were Black. Sex and age distribution for cases and controls were essentially identical between the full and present sample as well.

We found a positive association between $PM_{2.5}$ and PD where the RR for PD was 11%–28% higher in each of the 9 upper deciles of $PM_{2.5}$ relative to the lowest decile (Table 2). The association was strictly linear in the lower 5 deciles up to at least 13 $\mu g/m^3$ and then began to weaken as differences in $PM_{2.5}$ levels became more similar across deciles. Nonetheless, risk continued to increase generally, with exception of the highest 2 deciles, resulting in a significant trend overall ($p_{trend} < 0.0001$) (Table 2). When fitting a model with 2 linear splines, the model with the lowest (best) Akaike information criterion (814702.9) incorporated a knot at 13 $\mu g/m^3$ with 4.2% greater risk of PD for each additional $\mu g/m^3$ of $PM_{2.5}$ up to this level and then a weak, nonsignificant increase with increasing $PM_{2.5}$ levels thereafter resulting in a plateau (Table 2, overall $PM_{2.5}$-PD association $p < 0.0001$). The spline model was sensitive to knot placement, with 2 significant positive splines observed when the knot was at 12 $\mu g/m^3$ (RR = 1.048, 95% CI 1.041–1.055 per $\mu g/m^3$ of $PM_{2.5}$ up to this level and 1.007, 95% CI 1.003–1.012 per $\mu g/m^3$ of $PM_{2.5}$ thereafter, Akaike information criterion 814710.6), for example, but equally good yet contrasting models with a knot at either 14 or 15 $\mu g/m^3$. A simple linear model yielded a higher (poorer) Akaike information criterion (814770.9), and a likelihood ratio test confirmed a poorer fit relative to our best spline model ($p < 0.0001$). Based on the best spline model, the PD RR was 1.56 (95% CI 1.47–1.66) when comparing beneficiaries exposed to 12.93 $\mu g/m^3$ of $PM_{2.5}$ (median $PM_{2.5}$) with beneficiaries exposed to 2.16 $\mu g/m^3$ of $PM_{2.5}$ (minimum $PM_{2.5}$). With adjustment for 6 medical conditions associated with $PM_{2.5}$ that might facilitate PD diagnosis, the RR was 1.43 (95% CI 1.35–1.51). Adjustment for socioeconomic variables, access to neurologists, and pesticide use did not affect the RR for $PM_{2.5}$ and PD (eTable 1, links.lww.com/WNL/D163). Airborne trichloroethylene also did not confound the $PM_{2.5}$-PD

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**Table 2** Residential zip+4 Level $PM_{2.5}$ and PD Risk, US Medicare 2009

| $PM_{2.5}$ Decile | Zip+4 level $PM_{2.5}$, $\mu g/m^3$, range | RR (95% CI)[a] | Also adjusted for airborne trichloroethylene RR (95% CI)[a,b] |
|---|---|---|---|
| **D1** | 2.16 to 9.47 | 1.00 (Reference) | 1.00 (Reference) |
| **D2** | 9.50 to 10.87 | 1.11 (1.07–1.15)[c] | 1.11 (1.07–1.15)[c] |
| **D3** | 10.90 to 11.90 | 1.16 (1.12–1.20)[c] | 1.16 (1.12–1.20)[c] |
| **D4** | 11.93 to 12.90 | 1.22 (1.17–1.28)[c] | 1.21 (1.17–1.26)[c] |
| **D5** | 12.93 to 13.70 | 1.24 (1.19–1.28)[c] | 1.23 (1.19–1.28)[c] |
| **D6** | 13.73 to 14.27 | 1.23 (1.19–1.28)[c] | 1.23 (1.18–1.27)[c] |
| **D7** | 14.30 to 14.73 | 1.26 (1.22–1.31)[c] | 1.26 (1.21–1.30)[c] |
| **D8** | 14.77 to 15.33 | 1.28 (1.24–1.33)[c] | 1.28 (1.24–1.33)[c] |
| **D9** | 15.37 to 16.20 | 1.26 (1.21–1.30)[c] | 1.26 (1.21–1.30)[c] |
| **D10** | 16.23 to 23.30 | 1.24 (1.19–1.28)[c] | 1.23 (1.19–1.28)[c] |
| | | $p_{trend} < 0.0001$ | $p_{trend} < 0.0001$ |
| **$PM_{2.5}$ Spline** | | **RR (95% CI)[a] per $\mu g/m^3$ $PM_{2.5}$** | **RR (95% CI)[a] per $\mu g/m^3$ $PM_{2.5}$** |
| **Spline 1** | 2.16 to 13.00 | 1.042 (1.037–1.048)[c] | 1.042 (1.037–1.048)[c] |
| **Spline 2** | 13.03 to 23.30 | 1.002 (0.997–1.008) | 1.001 (0.996–1.007) |
| | | $p < 0.0001$[d] | $p < 0.0001$[d] |
| **Median vs minimum $PM_{2.5}$** | 12.93 vs 2.16 | 1.56 (1.47–1.66)[c] | 1.56 (1.47–1.66)[c] |

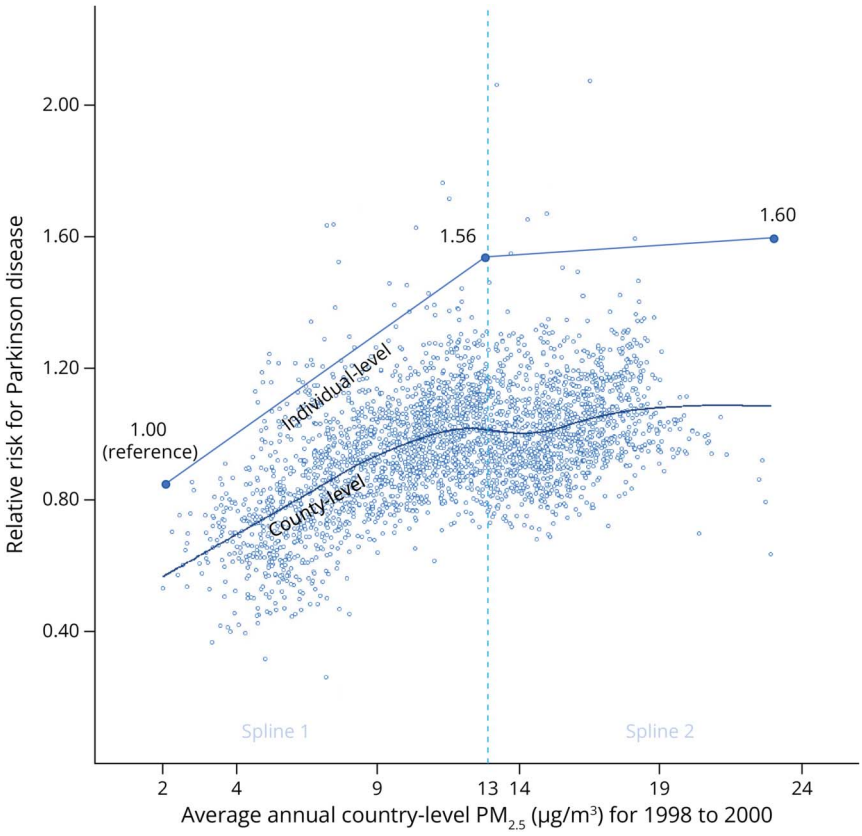
Abbreviations: $\mu g/m^3$ = =micrograms per cubic meter; CI = confidence interval; PD = Parkinson disease; $PM_{2.5}$ = particulate matter <2.5 micrometers in diameter; RR = relative risk.
[a] RRs estimated from odds ratios and associated 95% CIs from logistic regression with PD as the outcome and zip+4 level $PM_{2.5}$ as the independent variable, adjusted for individual-level variables (age, sex, race, ever/never smoking, and use of medical care), based on 65,180 cases and 15,561,435 controls with geocodable zip+4 data for residence 2 years before PD diagnosis or the control reference date and with $PM_{2.5}$ data at the geocoded location in the United States.
[b] Cases and controls were assigned the trichloroethylene value of the census tract that corresponded to their zip+4 center.
[c] $p < 0.001$.
[d] Overall $p$-value for the $PM_{2.5}$-PD association, assessed using a likelihood ratio test comparing models with and without both splines.

association (Table 2). In addition, the results were unchanged when Lewy body dementia was included in the case definition. We found no evidence that the $PM_{2.5}$-PD association differed by age ($p_{interaction} = 0.77$). The pattern of association was the same in men and women, but with slightly stronger associations in women for both splines ($p_{interaction} = 0.001$). The counties with average annual $PM_{2.5}$ levels that fell within Spline 1 were largely in the Western half of the United States (Figure 1). However, the results were fairly similar in urban areas, suburban areas, small towns, and rural areas, with significant RRs ranging from 1.040 to 1.047 per $\mu g/m^3$ of $PM_{2.5}$ up to 13 $\mu g/m^3$ and then relatively flat thereafter for each of land use type (likelihood ratio $p_{interaction} = 0.69$).

In univariate LISA (hot and cold spot) analysis, we found moderate clustering of similar values for PD risk across the contiguous United States (*Global Moran's I* value of 0.500, $p < 0.05$). The associated LISA map revealed an S-shaped pattern of high PD risk in the Mississippi-Ohio River Valley and a PD cold spot in a large portion of the Western part of the nation (Figure 2). PD risk was 19% greater in the Mississippi-Ohio River Valley hot spot compared with the rest of the nation. Bivariate LISAs revealed local spatial correlation characterized by high-high clusters (counties with above average PD risk and above average $PM_{2.5}$) and low-low clusters (counties with below average PD risk and below average $PM_{2.5}$) in the above respective areas (Figure 3). Several other smaller high-high or low-low areas and only a small number of discordant areas were observed. The Monte Carlo test of spatial variability further revealed spatial variation in the local parameter estimates for $PM_{2.5}$ as a PD risk factor ($p = 0.0001$). GWR identified specific clusters of counties where the positive association between $PM_{2.5}$ and PD risk was the strongest (Figure 4). The strongest positive coefficients formed a cluster of 51 counties within our PD cold spot in the mountainous regions of Colorado and Wyoming, where the $PM_{2.5}$-PD RR increased, in absolute terms, by approximately 0.15–0.16 per decile of $PM_{2.5}$ exposure. Although much weaker, we also identified a positive association between PD and $PM_{2.5}$ in our PD hot spot: For 118 counties North of the Mississippi-Ohio river confluence (in Indiana, Ohio, Illinois, and Missouri), the RR of PD increased by approximately 0.03–0.04 with each additional decile of

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**Figure 1** PM2.5 Exposure and Risk of PD Among US Medicare Beneficiaries in 2009



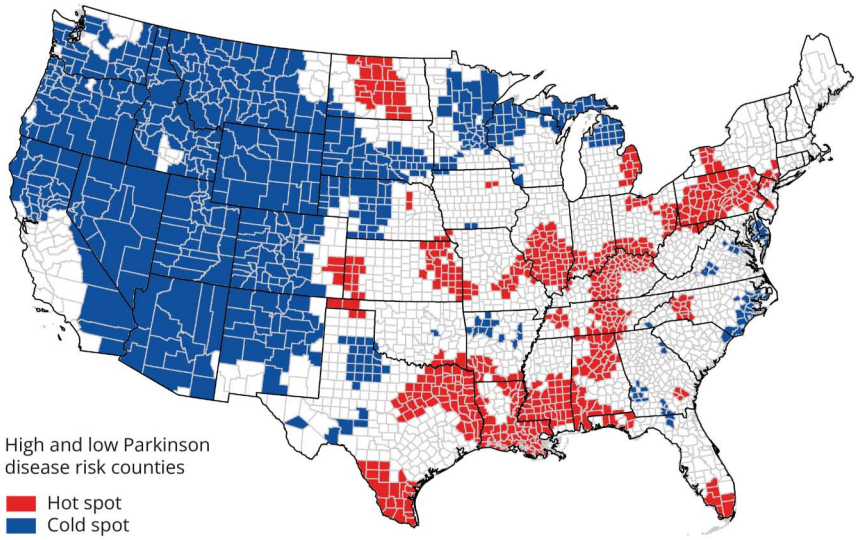

The relationship between individual-level (zip+4) $PM_{2.5}$ exposure and risk of PD among US Medicare beneficiaries in 2009 was well-described by a logistic regression model with 2 linear splines with a knot at 13 μg/m$^3$ $PM_{2.5}$. The adjusted relative risk of PD was 1.56 (95% CI 1.47–1.66) when comparing beneficiaries exposed to 12.93 μg/m$^3$ of $PM_{2.5}$ (the median of $PM_{2.5}$) with beneficiaries exposed to 2.16 μg/m$^3$ of $PM_{2.5}$ (the lowest level of $PM_{2.5}$ and the reference group for the individual-level line in the plot). The adjusted relative risk of PD was 1.60 (95% CI 1.51–1.70) when comparing beneficiaries exposed to the highest $PM_{2.5}$ with this same reference group. A locally weighted scatterplot smoothing curve between county-level mean $PM_{2.5}$ for the smoothed, standardized county-PD relative risk (with all other counties in the contiguous United States as the reference group). Each dot represents a county.

$PM_{2.5}$ exposure. Four regions of the country demonstrated null or inverse associations between $PM_{2.5}$ and PD. These included the North Dakota-Minnesota border; parts of the Mid-Atlantic; South Atlantic; and a region that spans part of Washington State, Idaho, and Montana (eFigure 1, links.lww.com/WNL/D162).

## Discussion

In this multimethod study investigating the nationwide patterns of PD risk and its relation to $PM_{2.5}$, we found an S-shaped pattern of high PD risk (hot spot) in the Mississippi-

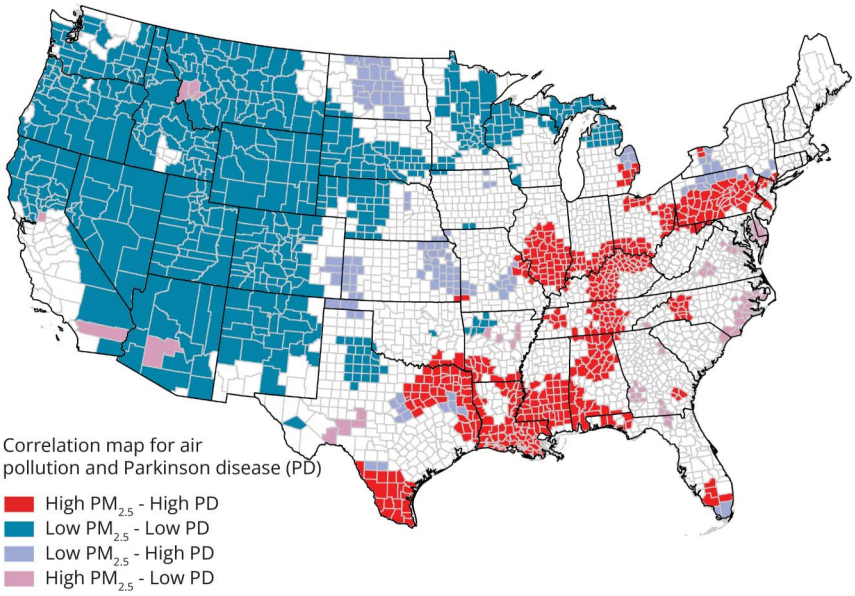
**Figure 2** Parkinson Disease Risk Hot and Cold Spots




Areas of high (hot spots) and low (cold spots) risk of PD among Medicare beneficiaries in the contiguous United States in 2009. Hot and cold spots were identified using univariate LISA analyses and county-PD relative risks that account for age, sex, race, smoking, health care utilization, and spatial dependency. LISA = local indicator of spatial association; PD = Parkinson disease.

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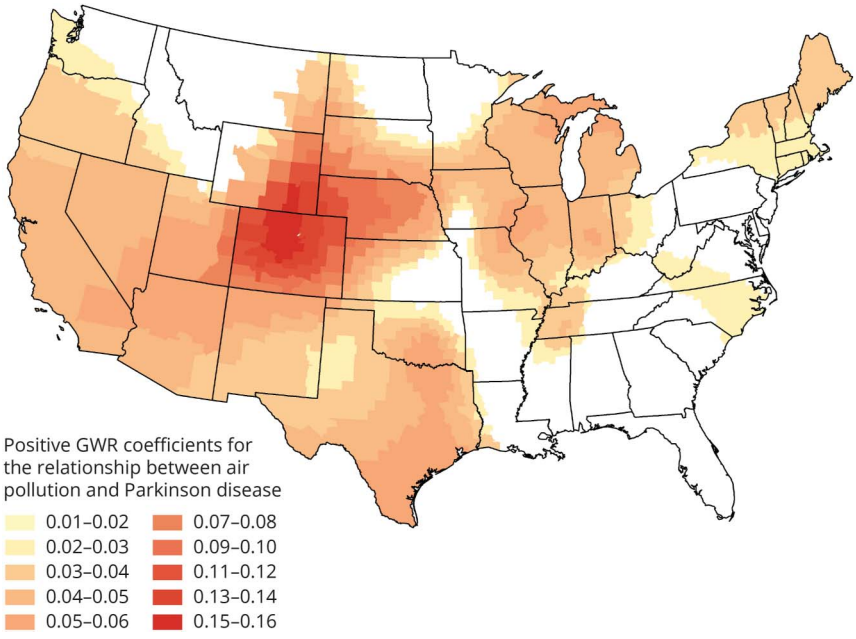
**Figure 3** Spatial Correlation Between $PM_{2.5}$ and Parkinson Disease Risk




Cluster map showing how PD hot and cold spots among US Medicare beneficiaries in 2009 overlay with $PM_{2.5}$ hot and cold spots in 1998–2000. Clusters were identified with bivariate LISA using (1) relative risks that account for age, sex, race, smoking, health care utilization, and spatial dependency and (2) average annual $PM_{2.5}$. The high-high category represents counties where risk of PD and $PM_{2.5}$ exposure are both high relative to means. The low-high and high-low clustering categories describe spatial outliers where a low-high county is one with below-average PD risk and above-average $PM_{2.5}$. The low-low category represents counties where PD risk and exposure are both less relative to their means. LISA = local indicator of spatial association; PD = Parkinson disease; $PM_{2.5}$ = particulate matter <2.5 micrometers in diameter.

Ohio River Valley and low PD risk (cold spot) in the Western United States. This geographic pattern aligns broadly with our prior study of incident PD[21] while providing a more refined pattern. Our bivariate LISA correlation map for PD and $PM_{2.5}$ closely overlapped with our PD risk map, in that PD hot spots generally aligned with $PM_{2.5}$ hot spots and PD cold spots generally aligned with $PM_{2.5}$ cold spots. Although bivariate LISA maps are exploratory and cannot alone confirm relationships at the nation level, we confirmed a $PM_{2.5}$-PD association nationally. In our regression analysis using individual-level data, the association between PD and $PM_{2.5}$ was linear up to at least 13 $\mu g/m^3$ $PM_{2.5}$. At the highest levels of $PM_{2.5}$, the relationship between $PM_{2.5}$ and PD appeared to plateau, but the overall association remained positive. Our GWR results also suggested a possible ceiling effect as the association weakened in the Mississippi-Ohio River Valley where some of the highest levels of $PM_{2.5}$ in the nation overlay regions with some of the highest PD risk in the nation.

**Figure 4** Strength of Positive Associations Between $PM_{2.5}$ and Parkinson Disease Risk




GWR coefficient map showing location and strength of significant positive coefficients for the association between county-level average annual $PM_{2.5}$ and Parkinson disease relative risk in the United States that account for age, sex, race, smoking, health care utilization, and spatial dependency. Each coefficient represents the absolute difference in the relative risk with each additional decile of $PM_{2.5}$ exposure. The strongest associations appear in Colorado and Wyoming, which have relatively low risk of Parkinson disease and relatively low $PM_{2.5}$ levels compared with the rest of the nation. This pattern of coefficients is an artifact of a possible ceiling effect in the Mississippi-Ohio River Valley, which has above-average Parkinson disease risk and some of the highest $PM_{2.5}$ levels in the nation. These patterns of coefficients are consistent with the observed positive association between $PM_{2.5}$ and PD nationwide overall that is stronger at lower $PM_{2.5}$ levels than higher $PM_{2.5}$ levels. GWR = geographically weighted regression; $PM_{2.5}$ = particulate matter <2.5 micrometers in diameter.

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Although the reason for the plateau is unclear, several studies report effect estimates consistent with a similar plateau.[12-14,16] It is more important that the robust dose-response association at the lower levels of $PM_{2.5}$ potentially has substantial public health relevance. Recently, the US Environmental Protection Agency proposed to revise the primary (health-based) annual $PM_{2.5}$ standard level from 12 μg/m$^3$ to 9–10 μg/m$^3$ because of growing evidence of health effects at levels lower than the previous regulatory standard.[44] Our study provides important additional evidence supporting this proposal. Previous studies,[8,16,20] but not all,[15] demonstrated clear linear associations at this lower range of $PM_{2.5}$ levels while the observed range of $PM_{2.5}$ levels or analytic approaches in all or most of the remainder could have obscured these associations. Therefore, our study provides insights that strengthen the interpretation of the broader literature as generally consistent with $PM_{2.5}$ as an exposure that increases risk of PD.

Key strengths of our study are that we used large, population-based data and were restricted to incident disease, aspects particularly important for geographic studies designed to inform disease etiology. Many prior studies used other PD-related outcomes that might be affected by PD progression or survival, not just risk. We chose to maximize sensitivity of identification of incident PD to ensure that our work was particularly resilient to selection bias.[26] Selection bias might occur in other studies with more restrictive definitions of incident PD if, for example, PD cases from higher vs lower $PM_{2.5}$ areas are less able to access neurologist care or survive long enough to begin anti-parkinsonian medications, that is, be ascertained as a case. In addition, our results are less vulnerable to biases from exposure measurement error because our zip+4 $PM_{2.5}$ data serves as a stronger proxy for individual-level exposure compared with environmental studies that use broader units of geography-based exposure assignment. Furthermore, our study leveraged new and innovative geographic information system methods. Geographic approaches offer insight beyond multilevel (cluster) regression modeling by enabling investigators to adjust for confounders within their maps and smooth extreme values by taking into account neighboring values, thus refining local patterns of disease. Another important strength is that, within the context of our population-based sample, we also aimed to be as inclusive as possible with not only case ascertainment but also study eligibility, applying only scientifically essential criteria. This allowed us to include minority, poor, and rural populations—and the PD cases that arise within these populations—that are often missed in epidemiologic studies of PD.[45]

We found a clear linear dose-response relation between $PM_{2.5}$ and PD up to at least 13 μg/m$^3$, and others have also observed a linear association across a similar range of $PM_{2.5}$ levels as in our study.[8,16,20] Two of these studies found a similar magnitude of effect, estimating ∼50% greater risk of PD when comparing ∼13 μg/m$^3$ of $PM_{2.5}$ with the lowest levels of exposure of ∼2 μg/m$^3$.[8,16] Within this approximate $PM_{2.5}$ range, the third, smaller study with local air monitoring data reported a linear association of ∼50% greater magnitude.[20] If the $PM_{2.5}$-PD association is causal, future studies investigating the association between PD and $PM_{2.5}$ subfractions might demonstrate an even greater effect size and might provide additional insights to PD etiology. A multicountry study in Europe confirms the importance of considering the subfractions of $PM_{2.5}$.[8]

Accordingly, one potential explanation for the heterogeneity in $PM_{2.5}$-PD associations across the many studies conducted to date[6-20] is that $PM_{2.5}$ is not a homogenous exposure, but rather a mixture of chemical components that varies by source. $PM_{2.5}$ can be produced from a variety of sources, including industrial emissions,[8] motor vehicle traffic,[19] and farming practices.[10] $PM_{2.5}$ may, therefore, have varied effects on the development of neurodegenerative disease depending on the chemical composition of $PM_{2.5}$ in different regions. Of particular relevance, studies have identified heavy metals within $PM_{2.5}$ in the Ohio River Valley, including manganese and zinc linked to iron and steel manufacturing.[46,47] In addition to the type of, and proximity to, emission sources, the physical environment may also influence levels of relevant $PM_{2.5}$ components and, hence, health effects. For example, the S-shaped pattern of high risk in the Mississippi-Ohio River Valley follows the low-elevation valleys along the Appalachian Mountains, suggesting a potential role of regional wind patterns and topography in the relationship between geographic exposures such as $PM_{2.5}$ and PD. Our methods allowed for the detection of differential effects across the subregions of the United States. In most regions, we detected a strong positive relationship between $PM_{2.5}$ and PD, whereas in a few regions, there was no detectable effect or there was evidence of inverse associations. These results might inform efforts to identify natural and built environment conditions that could provide some protections against air pollution, including specific climates and urban layouts. However, we cannot rule out the possibility that regional differences could be partly due to gene-environment interaction[48] and/or interactions with various nongenetic factors including diet.[49] The differential effects across regions emphasize the value of taking a geographic approach in large studies. Thus, we encourage the application of this approach in other parts of the world, including densely populated regions in South Asia where $PM_{2.5}$ levels can reach ∼60 μg/m$^3$ on average[50] (compared with 14 μg/m$^3$ for the contiguous United States at the time of our study).

There are several limitations of this work. Our geographic analyses relied on aggregate data subject to the ecological fallacy. Nevertheless, regression based on individual-level outcome data corroborated our county-level geographic findings. At the same time, we acknowledge that the separate issue of exposure measurement error (likely nondifferential) remains to some extent. Any $PM_{2.5}$ exposure model is imperfect, even the one we used, which achieved a cross-validated $R^2$ of 0.89.[28] In addition, we linked the objective

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$PM_{2.5}$ estimates from this model using zip+4 rather than exact addresses. Another limitation of Medicare data is the restriction on coverage to patients 65 years and older, although the $PM_{2.5}$-PD association did not differ by age. We excluded patients with Medicare Advantage plans from analyses because these plans do not report claims to Medicare, but during the study years, they only represented a quarter of beneficiaries,[25] so we anticipate that our results remain generalizable. We also excluded a small percentage of potential cases whose PD incident status was uncertain because of the diagnosis of Lewy body dementia and/or atypical parkinsonism. However, the results of our sensitivity analysis found no difference in the results when Lewy body dementia was included in our case definition. Our results also assume that beneficiaries were nonmobile during the 10 years before diagnosis; we acknowledge that our methods were unable to capture early-life exposures that might be relevant. That said, the overwhelming majority of beneficiaries in our study were nonmobile for at least the 5 years of data which were available to us. Owing to the long PD prodromal period,[25] we applied as much exposure lagging as possible and used $PM_{2.5}$ estimates from up to 10 years before diagnosis. The true period of relevant exposure may be longer, but we were unable to extend our exposure period or explore time windows of exposure given that earlier years of residence data were not available. We also acknowledge that the association between $PM_{2.5}$ and PD could result, at least in part, from a correlated, environmental exposure, but we ruled out confounding by both pesticides and trichloroethylene. In addition, while we found no evidence that the $PM_{2.5}$-PD association was stronger in men than women, we cannot rule out confounding by occupational exposures. Finally, we also note that the validity of our data on PD diagnosis requires individuals to obtain medical care and relies on competent diagnoses and data entry by health care providers and staff. A prior study suggested that the case identification method we used in our study had 82.7% sensitivity and 99.7% specificity among primary care patients,[26] and through our design and analysis choices, we sought to minimize the potential for selection bias and confounding by use of medical care. Despite limitations, our study advances current knowledge of $PM_{2.5}$ in relation to PD and demonstrates the utility of using a multimethod geographic approach for investigating environmental risk factors.

Using state-of-the-art geographic analytic techniques, we identified strong regional associations between $PM_{2.5}$ and PD in the United States. A deeper investigation into the subfractions of $PM_{2.5}$ in those regions may provide insight into PD risk factors.

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## References

1. Shaffer RM, Blanco MN, Li G, et al. Fine particulate matter and dementia incidence in the adult changes in thought study. *Environ Health Perspect.* 2021;129(8):87001. doi: 10.1289/EHP9018
2. Yuan S, Wang J, Jiang Q, et al. Long-term exposure to PM(2.5) and stroke: a systematic review and meta-analysis of cohort studies. *Environ Res.* 2019;177:108587. doi:10.1016/j.envres.2019.108587
3. Calderon-Garciduenas L, Solt AC, Henriquez-Roldan C, et al. Long-term air pollution exposure is associated with neuroinflammation, an altered innate immune response, disruption of the blood-brain barrier, ultrafine particulate deposition, and accumulation of amyloid beta-42 and alpha-synuclein in children and young adults. *Toxicol Pathol.* 2008;36(2):289-310. doi:10.1177/0192623307313011

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4. Lee CP, Zhu CH, Su CC. Increased prevalence of Parkinson's disease in soils with high arsenic levels. *Parkinsonism Relat Disord.* 2021;88:19-23. doi:10.1016/j.parkreldis.2021.05.029
5. Willis AW, Evanoff BA, Lian M, et al. Metal emissions and urban incident Parkinson disease: a community health study of medicare beneficiaries by using geographic information systems. *Am J Epidemiol.* 2010;172(12):1357-1363. doi:10.1093/aje/kwq303
6. Cerza F, Renzi M, Agabiti N, et al. Residential exposure to air pollution and incidence of Parkinson's disease in a large metropolitan cohort. *Environ Epidemiol.* 2018;2(3):e023. doi:10.1097/ee9.0000000000000023
7. Chen H, Kwong JC, Copes R, et al. Living near major roads and the incidence of dementia, Parkinson's disease, and multiple sclerosis: a population-based cohort study. *Lancet.* 2017;389(10070):718-726. doi:10.1016/S0140-6736(16)32399-6
8. Cole-Hunter T, Zhang J, So R, et al. Long-term air pollution exposure and Parkinson's disease mortality in a large pooled European cohort: an ELAPSE study. *Environ Int.* 2023;171:107667. doi:10.1016/j.envint.2022.107667
9. Jo S, Kim YJ, Park KW, et al. Association of $NO_2$ and other air pollution exposures with the risk of Parkinson disease. *JAMA Neurol.* 2021;78(7):800-808. doi:10.1001/jamaneurol.2021.1335
10. Kirrane EF, Bowman C, Davis JA, et al. Associations of ozone and $PM_{2.5}$ concentrations with Parkinson's disease among participants in the Agricultural Health Study. *J Occup Environ Med.* 2015;57(5):509-517. doi:10.1097/JOM.0000000000000451
11. Kioumourtzoglou MA, Schwartz JD, Weisskopf MG, et al. Long-term $PM_{2.5}$ exposure and neurological hospital admissions in the Northeastern United States. *Environ Health Perspect.* 2016;124(1):23-29. doi:10.1289/ehp.1408973
12. Liu R, Young MT, Chen JC, Kaufman JD, Chen H. Ambient air pollution exposures and risk of Parkinson disease. *Environ Health Perspect.* 2016;124(11):1759-1765. doi:10.1289/EHP135
13. Palacios N, Fitzgerald KC, Hart JE, et al. Particulate matter and risk of Parkinson disease in a large prospective study of women. *Environ Health.* 2014;13:80. doi:10.1186/1476-069X-13-80
14. Palacios N, Fitzgerald KC, Hart JE, et al. Air pollution and risk of Parkinson's disease in a large prospective study of men. *Environ Health Perspect.* 2017;125(8):087011. doi:10.1289/EHP259
15. Salimi F, Hanigan I, Jalaludin B, et al. Associations between long-term exposure to ambient air pollution and Parkinson's disease prevalence: a cross-sectional study. *Neurochem Int.* 2020;133:104615. doi:10.1016/j.neuint.2019.104615
16. Shi L, Wu X, Danesh Yazdi M, et al. Long-term effects of PM(2.5) on neurological disorders in the American Medicare population: a longitudinal cohort study. *Lancet Planet Health.* 2020;4(12):e557-e565. doi:10.1016/S2542-5196(20)30227-8
17. Shin S, Burnett RT, Kwong JC, et al. Effects of ambient air pollution on incident Parkinson's disease in Ontario, 2001 to 2013: a population-based cohort study. *Int J Epidemiol.* 2018;47(6):2038-2048. doi:10.1093/ije/dyy172
18. Toro R, Downward GS, van der Mark M, et al. Parkinson's disease and long-term exposure to outdoor air pollution: a matched case-control study in the Netherlands. *Environ Int.* 2019;129:28-34. doi:10.1016/j.envint.2019.04.069
19. Yu Z, Wei F, Zhang X, et al. Air pollution, surrounding green, road proximity and Parkinson's disease: a prospective cohort study. *Environ Res.* 2021;197:111170. doi:10.1016/j.envres.2021.111170
20. Yuchi W, Sbihi H, Davies H, Tamburic L, Brauer M. Road proximity, air pollution, noise, green space and neurologic disease incidence: a population-based cohort study. *Environ Health.* 2020;19(1):8. doi:10.1186/s12940-020-0565-4
21. Wright Willis A, Evanoff BA, Lian M, Criswell SR, Racette BA. Geographic and ethnic variation in Parkinson disease: a population-based study of US Medicare beneficiaries. *Neuroepidemiology.* 2010;34(3):143-151. doi:10.1159/000275491
22. Lux WE, Kurtzke JF. Is Parkinson's disease acquired? Evidence from a geographic comparison with multiple sclerosis. *Neurology.* 1987;37(3):467-471. doi:10.1212/wnl.37.3.467
23. Kurtzke JF, Goldberg ID. Parkinsonism death rates by race, sex, and geography. *Neurology.* 1988;38(10):1558-1561. doi:10.1212/wnl.38.10.1558
24. Lanska DJ. The geographic distribution of Parkinson's disease mortality in the United States. *J Neurol Sci.* 1997;150(1):63-70. doi:10.1016/s0022-510x(97)05371-9
25. Searles Nielsen S, Warden MN, Camacho-Soto A, Willis AW, Wright BA, Racette BA. A predictive model to identify Parkinson disease from administrative claims data. *Neurology.* 2017;89(14):1448-1456. doi:10.1212/WNL.0000000000004536
26. Butt DA, Tu K, Young J, et al. A validation study of administrative data algorithms to identify patients with Parkinsonism with prevalence and incidence trends. *Neuroepidemiology.* 2014;43(1):28-37. doi:10.1159/000365590
27. Di Q, Wei Y, Shtein A, et al. *Daily and Annual $PM_{2.5}$ Concentrations for the Contiguous United States, 1-km Grids, v1 (2000-2016)*; 2021. Palisades, New York: NASA Socioeconomic Data and Applications Center (SEDAC). Accessed October 10, 2023. https://sedac.ciesin.columbia.edu/data/set/aqdh-pm2-5-concentrations-contiguous-us-1-km-2000-2016
28. Di Q, Amini H, Shi L, et al. An ensemble-based model of PM(2.5) concentration across the contiguous United States with high spatiotemporal resolution. *Environ Int.* 2019;130:104909. doi:10.1016/j.envint.2019.104909
29. Dwyer-Lindgren L, Mokdad AH, Srebotnjak T, Flaxman AD, Hansen GM, Murray CJ. Cigarette smoking prevalence in US counties: 1996-2012. *Popul Health Metr.* 2014;12(1):5. doi:10.1186/1478-7954-12-5
30. Faust I, Warden M, Camacho-Soto A, Racette BA, Searles Nielsen S. A predictive algorithm to identify ever smoking in medical claims-based epidemiologic studies. *Ann Epidemiol.* 2023;85:59-67.e6. doi:10.1016/j.annepidem.2023.04.019
31. U.S. Environmental Protection Agency. *National Air Toxics Assessment (NATA): 2005 Census Tract Modeled Ambient Concentrations, Exposures, and Risks*; 2005. Accessed March 4, 2023 epa.gov/national-air-toxics-assessment/2005-nata-assessment-results.
32. Kind AJH, Buckingham WR. Making neighborhood-disadvantage metrics accessible - the neighborhood Atlas. *N Engl J Med.* 2018;378(26):2456-2458. doi:10.1056/NEJMp1802313
33. U.S. Census Bureau. *2009-2011 American Community Survey 3-year Public Use Microdata Samples [SAS Data file]*; 2010. Accessed October 20, 2022. factfinder.census.gov/faces/nav/jsf/pages/searchresults.xhtml?refresh=t
34. U.S. Geological Survey (USGS). *Estimated Annual Agricultural Pesticide Use: County-Level Data. National Water-Quality Assessment (NAWQA) Project; Pesticide National Synthesis Project*. U.S. Department of the Interior; 2018. Accessed October 18, 2019. water.usgs.gov/nawqa/pnsp/usage/maps/county-level/.
35. Rue H, Martino S, Chopin N. Approximate Bayesian inference for latent Gaussian models by using integrated nested Laplace approximations. *J R Stat Soc Ser B: Stat Methodol.* 2009;71(2):319-392. doi:10.1111/j.1467-9868.2008.00700.x
36. Reeve A, Simcox E, Turnbull D. Ageing and Parkinson's disease: why is advancing age the biggest risk factor? *Ageing Res Rev.* 2014;14(100):19-30. doi:10.1016/j.arr.2014.01.004
37. Gillies GE, Pienaar IS, Vohra S, Qamhawi Z. Sex differences in Parkinson's disease. *Front Neuroendocrinol.* 2014;35(3):370-384. doi:10.1016/j.yfrne.2014.02.002
38. Branson CO, Ferree A, Hohler AD, Saint-Hilaire M. Racial disparities in Parkinson disease: a systematic review of the literature. *Adv Parkinson's Dis.* 2016;05(4):87-96. doi:10.4236/apd.2016.54011
39. Nefzger MD, Quadfasel FA, Karl VC. A retrospective study of smoking in Parkinson's disease. *Am J Epidemiol.* 1968;88(2):149-158. doi:10.1093/oxfordjournals.aje.a120874
40. Gross A, Racette BA, Camacho-Soto A, Dube U, Searles Nielsen S. Use of medical care biases associations between Parkinson disease and other medical conditions. *Neurology.* 2018;90(24):e2155–e2165. doi:10.1212/WNL.0000000000005678
41. Anselin L, Syabri I, Kho Y. GeoDa: an introduction to spatial data analysis. *Geogr Anal.* 2006;38(1):5-22. doi:10.1111/j.0016-7363.2005.00671.x
42. Centers for Medicare and Medicaid Services (CMS). *Research Identifiable Files (RIFs)*. Centers for Medicare and Medicaid Services (CMS); 2009.
43. Silver MR, Racette BA, Dube U, Faust IM, Searles Nielsen S. Well water and Parkinson's disease in medicare beneficiaries: a Nationwide Case-Control Study. *J Parkinsons Dis.* 2020;10(2):693-705. doi:10.3233/JPD-191793
44. U.S. Environmental Protection Agency. *National Ambient Air Quality Standards (NAAQS) for PM: Rule Summary*; 2023. Accessed February 16, 2023 epa.gov/pm-pollution/national-ambient-air-quality-standards-naaqs-pm#rule-summary.
45. Willis AW, Schootman M, Evanoff BA, Perlmutter JS, Racette BA. Neurologist care in Parkinson disease: a utilization, outcomes, and survival study. *Neurology.* 2011;77(9):851-857. doi:10.1212/WNL.0b013e31822c9123
46. Vedantham R, Landis MS, Olson D, Pancras JP. Source identification of PM2.5 in Steubenville, Ohio using a hybrid method for highly time-resolved data. *Environ Sci Technol.* 2014;48(3):1718-1726. doi:10.1021/es402704n
47. Kim M, Deshpande SR, Crist KC. Source apportionment of fine particulate matter ($PM_{2.5}$) at a rural Ohio River Valley site. *Atmos Environ.* 2007;41(39):9231-9243. doi:10.1016/j.atmosenv.2007.07.061
48. Lee PC, Raaschou-Nielsen O, Lill CM, et al. Gene-environment interactions linking air pollution and inflammation in Parkinson's disease. *Environ Res.* 2016;151:713-720. doi:10.1016/j.envres.2016.09.006
49. Chen C, Hayden KM, Kaufman JD, et al. Adherence to a MIND-like dietary pattern, long-term exposure to fine particulate matter air pollution, and MRI-based measures of brain volume: the Women's Health Initiative Memory Study-MRI. *Environ Health Perspect.* 2021;129(12):127008. doi:10.1289/EHP8036
50. van Donkelaar A, Martin RV, Brauer M, et al. Global estimates of fine particulate matter using a combined geophysical-statistical method with information from satellites, models, and monitors. *Environ Sci Technol.* 2016;50(7):3762-3772. doi:10.1021/acs.est.5b05833


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